

UNIVERSITY OF MASSACHUSETTS  
Department of Mathematics and Statistics  
Advanced Exam - Probability and Mathematical Statistics  
Wednesday, August 28, 2013

Seventy points are required to pass, and at least twenty-five must come from both the probability section (problems 1-3) and the multivariate / linear models section (problems 4-5).

1. Assume that  $\phi(t)$  is the characteristic function of a random variable  $X$ . (20 points)
  - (a) Prove that  $|\phi(t)|^2$  is also the characteristic function of a random variable.
  - (b) Assume that  $\phi'(t)$  exists for all  $t$  in some neighborhood of 0, and

$$\lim_{t \rightarrow 0} \frac{\phi(t) - 1}{t^2} = \frac{1}{2}\sigma^2 < \infty$$

Prove that  $E(X) = 0$  and  $E(X^2) = \sigma^2$ .

2. Let  $(\Omega, \mathcal{F}, P)$  be a probability space. Assume all random variables we consider here have finite expectations. Let  $\mathcal{D} = (D_1, D_2, \dots)$  be the countable partition of  $\Omega$ , such that  $P(D_n) > 0$ , for any  $n \geq 1$ ,  $D_n \cap D_m = \emptyset$ , for any  $m \neq n$ . Let  $\mathcal{G} = \sigma(\mathcal{D})$  be the smallest  $\sigma$ -algebra generated by  $\mathcal{D}$ . Consider the following questions. (25 points)
  - (a) If  $X \in \mathcal{F}$ , write the explicit expression of  $E(X|\mathcal{G})$ .
  - (b) If  $Y \in \mathcal{D}$ , what is  $E(XY|\sigma(\mathcal{D}))$ ?
  - (c) Let  $\mathcal{D}_1$  be another countable partition of  $\Omega$  defined similar as  $\mathcal{D}$ , such that  $\mathcal{D}_1$  is coarser than  $\mathcal{D}$ . i.e. for any  $D \in \mathcal{D}_1$ , there exists  $A_1, A_2, \dots, A_n$  in  $\mathcal{D}$ , such that  $D = A_1 \cup A_2 \cup \dots \cup A_n$ . Find  $E(E(X|\sigma(\mathcal{D}))|\sigma(\mathcal{D}_1))$  and  $E(E(X|\mathcal{D}_1)|\sigma(\mathcal{D}))$ ?
  - (d) Show that if  $X$  and  $Y$  are random variables with  $E(Y|\mathcal{G}) = X$  and  $E(Y^2) = E(X^2) < \infty$ , then  $X = Y$  a.s.
3. Prove that the sum of two independent normal random variables is a normal variable. More precisely, if  $X \sim N(\mu_1, \sigma_1^2)$ ,  $Y \sim N(\mu_2, \sigma_2^2)$  are independent, then  $X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$ . (5 points)
4. Let  $Y$  and  $X$  both be random variables. A linear predictor of  $Y$  from  $X$  is a function of the form  $\beta_0 + \beta_1 X$ , and we define the best linear predictor to be one that minimizes  $E\{Y - (\beta_0 + \beta_1 X)\}^2$ . (25 points)
  - (a) Derive the best linear predictor of  $Y$  from  $X$ .
  - (b) Let  $\hat{Y}^{LIN}$  be the best linear predictor. Show that it unbiased for  $Y$ .
  - (c) Derive the squared prediction error,  $E(Y - \hat{Y}^{LIN})^2$ .

- (d) Suppose that the random vector  $(Y, X)$  has a multivariate normal distribution with mean  $\mu_y, \mu_x$  and covariance matrix  $\begin{pmatrix} \sigma_y^2 & \sigma_{yx} \\ \sigma_{yx} & \sigma_x^2 \end{pmatrix}$ . How do your answers above relate to this distribution?
5. Suppose  $Y = X\beta + \varepsilon$  when  $Y$  is a random vector of length  $n$ ,  $X$  is  $n$  by  $p$  matrix with rank  $p$ ,  $\beta$  is a vector of length  $p$ , and  $\varepsilon$  is a random  $n$  vector with independent  $N(0, \sigma^2)$  components. (25 points)
- (a) Derive the maximum likelihood estimators of  $\beta$  and  $\sigma^2$ .
  - (b) Let  $\hat{\beta}$  be your estimator for  $\beta$  from part (a). Derive its sampling distribution.
  - (c) Let  $\hat{\sigma}^2$  be your estimator for  $\sigma^2$  from part (a). Derive the sampling distribution of  $n\hat{\sigma}^2/\sigma^2$ .
  - (d) Show that  $\hat{Y} = X\hat{\beta}$  and  $R = Y - \hat{Y}$  are independent.
  - (e) Are  $\hat{\beta}$  and  $\hat{\sigma}^2$  independent? Why or why not?